

Principal Components

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Outline

- Introduction
- Popular Principal Components
- Summarizing Sample Variation by Principal Components
- Graphing the Principal Components
- Large Sample Inferences
- Monitoring Quality with Principal Components

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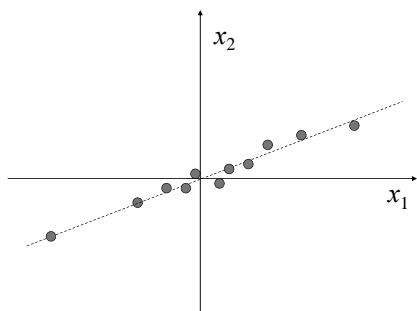
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Questions

- What is the concept of the Principal Components?
- What are the objectives of the Principal Components?

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Concept of Principal Components



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Principal Component Analysis

- Explain the variance-covariance structure of a set of variables through a few linear combinations of these variables
- Objectives
 - Data reduction
 - Interpretation
- Does not need normality assumption in general

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Questions

- How to find the Principal Components for a Random vector with a known probability distribution? (Result 8.1)
- What is the relationship between the sum of all eigenvalues and the trace of the covariance matrix? (Result 8.2)
- How to calculate the proportion of total population variance due to the k th principal component?

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Questions

- What is the relationship between the i th principal component and the k th variable? (Result 8.3)
- What is the geometric interpretation of the principal components?
- How to find the principal components for a standardized random vector? (Result 8.4)

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Questions

- What are the principal components for a diagonal covariance matrix?
- What are the principal components for the special covariance matrix

$$\Sigma = \begin{bmatrix} \sigma^2 & \rho\sigma^2 & \dots & \rho\sigma^2 \\ \rho\sigma^2 & \sigma^2 & \dots & \rho\sigma^2 \\ \vdots & \vdots & \ddots & \vdots \\ \rho\sigma^2 & \rho\sigma^2 & \dots & \sigma^2 \end{bmatrix}$$

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Principal Components

Random vector $\mathbf{X}' = [X_1 \quad X_2 \quad \dots \quad X_p]$ has the covariance matrix Σ

Linear combination: $Y_i = \mathbf{a}_i' \mathbf{X}$, $i = 1, 2, \dots, p$

$\text{Var}(Y_i) = \mathbf{a}_i' \Sigma \mathbf{a}_i$, $\text{Cov}(Y_i, Y_k) = \mathbf{a}_i' \Sigma \mathbf{a}_k$

First principal component:

$\mathbf{a}_1' \mathbf{X}$ that maximizes $\text{Var}(\mathbf{a}_1' \mathbf{X})$ subject to $\mathbf{a}_1' \mathbf{a}_1 = 1$

i th principal component:

$\mathbf{a}_i' \mathbf{X}$ that maximizes $\text{Var}(\mathbf{a}_i' \mathbf{X})$ subject to $\mathbf{a}_i' \mathbf{a}_i = 1$

and $\text{Cov}(\mathbf{a}_i' \mathbf{X}, \mathbf{a}_k' \mathbf{X}) = 0$ for $k < i$

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Result 8.1

Covariance matrix Σ of random vector \mathbf{X} is with eigenvalue-eigenvector pairs $(\lambda_i, \mathbf{e}_i)$, where $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$

The i th principal component is given by

$Y_i = \mathbf{e}_i' \mathbf{X}$, $i = 1, 2, \dots, p$, with

$\text{Var}(Y_i) = \mathbf{e}_i' \Sigma \mathbf{e}_i = \lambda_i$, $i = 1, 2, \dots, p$

$\text{Cov}(Y_i, Y_k) = \mathbf{e}_i' \Sigma \mathbf{e}_k = 0$, $i \neq k$

If some λ_i are equal, the choice of corresponding \mathbf{e}_i and hence Y_i are not unique

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Proof of Result 8.1

$$\begin{aligned} \max_{\mathbf{a} \neq 0} \frac{\mathbf{a}' \Sigma \mathbf{a}}{\mathbf{a}' \mathbf{a}} &= \lambda_1 \text{ attained when } \mathbf{a} = \mathbf{e}_1 \\ \mathbf{e}_1' \mathbf{e}_1 &= 1, \text{ thus } \max_{\mathbf{a} \neq 0} \frac{\mathbf{a}' \Sigma \mathbf{a}}{\mathbf{a}' \mathbf{a}} = \lambda_1 = \mathbf{e}_1' \Sigma \mathbf{e}_1 = \text{Var}(Y_1) \\ \max_{\mathbf{a} \perp \mathbf{e}_1, \dots, \mathbf{e}_k} \frac{\mathbf{a}' \Sigma \mathbf{a}}{\mathbf{a}' \mathbf{a}} &= \lambda_{k+1}, k = 1, 2, \dots, p-1 \\ \mathbf{a} = \mathbf{e}_{k+1}, \mathbf{e}_{k+1}' \Sigma \mathbf{e}_{k+1} &= \lambda_{k+1} = \text{Var}(Y_{k+1}) \\ \text{Cov}(Y_i, Y_k) = \mathbf{e}_i' \Sigma \mathbf{e}_k &= \mathbf{e}_i' \lambda_k \mathbf{e}_k = 0 \text{ for any } i \neq k \end{aligned}$$

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Result 8.2

Covariance matrix Σ of random vector $\mathbf{X} = [X_1 \ X_2 \ \dots \ X_p]$ is with eigenvalue-eigenvector pairs $(\lambda_i, \mathbf{e}_i)$, where $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$. The i th principal component is given by $Y_i = \mathbf{e}_i' \mathbf{X}$, $i = 1, 2, \dots, p$, then

$$\begin{aligned} \sigma_{11} + \sigma_{22} + \dots + \sigma_{pp} &= \sum_{i=1}^p \text{Var}(X_i) \\ &= \lambda_1 + \lambda_2 + \dots + \lambda_p = \sum_{i=1}^p \text{Var}(Y_i) \end{aligned}$$

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Proof of Result 8.2

$$\begin{aligned} \Sigma &= \mathbf{P} \Lambda \mathbf{P}', \quad \Lambda = \text{diag}\{\lambda_1, \lambda_2, \dots, \lambda_p\} \\ \mathbf{P} &= [\mathbf{e}_1 \ \mathbf{e}_2 \ \dots \ \mathbf{e}_p], \quad \mathbf{P} \mathbf{P}' = \mathbf{P}' \mathbf{P} = \mathbf{I} \\ \sigma_{11} + \sigma_{22} + \dots + \sigma_{pp} &= \sum_{i=1}^p \text{Var}(X_i) = \text{tr}(\Sigma) \\ &= \text{tr}(\mathbf{P} \Lambda \mathbf{P}') = \text{tr}(\Lambda \mathbf{P}' \mathbf{P}) = \text{tr}(\Lambda) \\ &= \lambda_1 + \lambda_2 + \dots + \lambda_p = \sum_{i=1}^p \text{Var}(Y_i) \end{aligned}$$

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Proportion of Total Variance due to the k th Principal Component

$$\left(\begin{array}{l} \text{Proportion of total} \\ \text{population variance} \\ \text{due to the } k\text{th principal} \\ \text{component} \end{array} \right) = \frac{\lambda_k}{\lambda_1 + \lambda_2 + \dots + \lambda_p}$$

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Result 8.3

$Y_i = \mathbf{e}_i' \mathbf{X}$ are the principal components obtained from the covariance matrix Σ , then

$$\rho_{Y_i, X_k} = \frac{e_{ik} \sqrt{\lambda_i}}{\sqrt{\sigma_{kk}}}, \quad i, k = 1, 2, \dots, p$$

are the correlation coefficients between Y_i and variable X_k . Here $\mathbf{e}_i' = [e_{i1} \ e_{i2} \ \dots \ e_{ip}]$ is the eigenvector of Σ corresponding to the eigenvalue λ_i . Also, $\mathbf{X}' = [X_1 \ X_2 \ \dots \ X_p]$

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Proof of Result 8.3

$\mathbf{a}_k' = [0 \ \dots \ 0 \ 1 \ 0 \ \dots \ 0]$ so that $X_k = \mathbf{a}_k' \mathbf{X}$

$$\text{Cov}(X_k, Y_i) = \text{Cov}(\mathbf{a}_k' \mathbf{X}, \mathbf{e}_i' \mathbf{X}) = \mathbf{a}_k' \Sigma \mathbf{e}_i = \lambda_i e_{ik}$$

$$\text{Var}(Y_i) = \lambda_i, \quad \text{Var}(X_k) = \sigma_{kk}$$

$$\begin{aligned} \rho_{Y_i, X_k} &= \frac{\text{Cov}(X_k, Y_i)}{\sqrt{\text{Var}(Y_i)} \sqrt{\text{Var}(X_k)}} = \frac{\lambda_i e_{ik}}{\sqrt{\lambda_i} \sqrt{\sigma_{kk}}} \\ &= \frac{\sqrt{\lambda_i} e_{ik}}{\sqrt{\sigma_{kk}}} \quad i, k = 1, 2, \dots, p \end{aligned}$$

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Example 8.1

$\mathbf{X}' = [X_1 \ X_2 \ X_3]$ has the covariance matrix
 $\Sigma = \begin{bmatrix} 1 & -2 & 0 \\ -2 & 5 & 0 \\ 0 & 0 & 2 \end{bmatrix}$, whose eigenvalue - eigenvector

pairs are

$$\lambda_1 = 5.83, \quad \mathbf{e}_1' = [0.383 \quad -0.924 \quad 0]$$

$$\lambda_2 = 2.00, \quad \mathbf{e}_2' = [0 \quad 0 \quad 1]$$

$$\lambda_3 = 0.17, \quad \mathbf{e}_3' = [0.924 \quad 0.383 \quad 0]$$

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Example 8.1

Principal components

$$Y_1 = \mathbf{e}_1' \mathbf{X} = 0.383X_1 - 0.924X_2$$

$$Y_2 = \mathbf{e}_2' \mathbf{X} = X_3$$

$$Y_3 = \mathbf{e}_3' \mathbf{X} = 0.924X_1 + 0.383X_2$$

Verification

$$\text{Var}(Y_1) = (0.383)^2 \text{Var}(X_1)$$

$$+ 2(0.383)(-0.924) \text{Cov}(X_1, X_2) + (-0.924)^2 \text{Var}(X_2)$$

$$= 5.83 = \lambda_1$$

$$\text{Cov}(Y_1, Y_2) = 0.383 \text{Cov}(X_1, X_3) - 0.924 \text{Cov}(X_2, X_3) = 0$$

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Example 8.1

$$\sigma_{11} + \sigma_{22} + \sigma_{33} = 8 = 5.83 + 2.00 + 0.17 = \lambda_1 + \lambda_2 + \lambda_3$$

$$\frac{\lambda_1}{\lambda_1 + \lambda_2 + \lambda_3} = 0.73, \quad \frac{\lambda_1 + \lambda_2}{\lambda_1 + \lambda_2 + \lambda_3} = 0.98$$

$$\rho_{Y_1, X_1} = \frac{e_{11}\sqrt{\lambda_1}}{\sqrt{\sigma_{11}}} = \frac{0.383\sqrt{0.583}}{\sqrt{1}} = 0.925$$

$$\rho_{Y_1, X_2} = \frac{e_{12}\sqrt{\lambda_1}}{\sqrt{\sigma_{22}}} = \frac{-0.924\sqrt{0.583}}{\sqrt{5}} = -0.998$$

$$\rho_{Y_2, X_1} = \rho_{Y_2, X_2} = 0, \quad \rho_{Y_2, X_3} = \frac{\sqrt{\lambda_2}}{\sqrt{\sigma_{33}}} = 1$$

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Geometrical Interpretation

$\mathbf{X} : N_p(\boldsymbol{\mu}, \Sigma)$

Σ is with eigenvalue - eigenvector pairs $(\lambda_i, \mathbf{e}_i)$

constant probability density ellipsoid

$$(\mathbf{x} - \boldsymbol{\mu})' \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) = c^2$$

$$c^2 = \frac{1}{\lambda_1} (\mathbf{e}_1' (\mathbf{x} - \boldsymbol{\mu}))^2 + \frac{1}{\lambda_2} (\mathbf{e}_2' (\mathbf{x} - \boldsymbol{\mu}))^2 + \dots + \frac{1}{\lambda_p} (\mathbf{e}_p' (\mathbf{x} - \boldsymbol{\mu}))^2$$

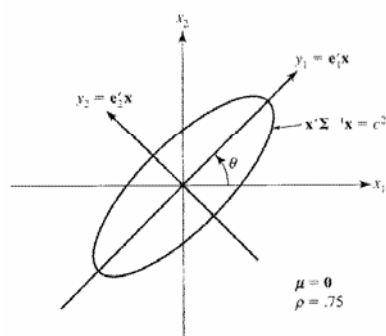
Principal components of $\mathbf{x} - \boldsymbol{\mu} : y_i = \mathbf{e}_i' (\mathbf{x} - \boldsymbol{\mu})$

$i = 1, 2, \dots, p$

$$c^2 = \frac{1}{\lambda_1} y_1^2 + \frac{1}{\lambda_2} y_2^2 + \dots + \frac{1}{\lambda_p} y_p^2$$

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Geometric Interpretation



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Standardized Variables

$$Z_i = \frac{X_i - \mu_i}{\sqrt{\sigma_{ii}}}, \quad i = 1, 2, \dots, p$$

$$\mathbf{Z} = \mathbf{V}^{-1/2} (\mathbf{X} - \boldsymbol{\mu}), \quad \mathbf{V}^{1/2} = \begin{bmatrix} \sqrt{\sigma_{11}} & 0 & \dots & 0 \\ 0 & \sqrt{\sigma_{22}} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \sqrt{\sigma_{pp}} \end{bmatrix}$$

$$\text{Cov}(\mathbf{Z}) = \mathbf{V}^{-1/2} \Sigma \mathbf{V}^{1/2} = \boldsymbol{\rho} = \begin{bmatrix} 1 & \rho_{12} & \dots & \rho_{1p} \\ \rho_{12} & 1 & \dots & \rho_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \rho_{1p} & \rho_{2p} & \dots & 1 \end{bmatrix}$$

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Result 8.4

$\mathbf{Z}' = [Z_1 \ Z_2 \ \dots \ Z_p]$ with $\text{Cov}(\mathbf{Z}) = \boldsymbol{\rho}$

$(\lambda_i, \mathbf{e}_i)$: eigenvalue - eigenvector pairs of $\boldsymbol{\rho}$

$$\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$$

The i th principal component of \mathbf{Z} :

$$Y_i = \mathbf{e}_i' \mathbf{Z} = \mathbf{e}_i' \mathbf{V}^{-1/2} (\mathbf{X} - \boldsymbol{\mu}), \quad i = 1, 2, \dots, p$$

$$\sum_{i=1}^p \text{Var}(Y_i) = \sum_{i=1}^p \text{Var}(Z_i) = p$$

$$\rho_{Y_i, Z_k} = e_{ik} \sqrt{\lambda_i}, \quad i, k = 1, 2, \dots, p$$

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Proportion of Total Variance due to the k th Principal Component

$$\left(\begin{array}{l} \text{Proportion of (standardized)} \\ \text{population variance} \\ \text{due to the } k\text{th principal} \\ \text{component} \end{array} \right) = \frac{\lambda_k}{p}, \quad k = 1, 2, \dots, p$$

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Example 8.2

$$\boldsymbol{\Sigma} = \begin{bmatrix} 1 & 4 \\ 4 & 100 \end{bmatrix}, \quad \boldsymbol{\rho} = \begin{bmatrix} 1 & 0.4 \\ 0.4 & 1 \end{bmatrix}$$

Eigenvalue - eigenvector pairs for $\boldsymbol{\Sigma}$:

$$\lambda_1 = 100.16, \quad \mathbf{e}_1' = [0.040 \quad 0.999]$$

$$\lambda_2 = 0.84, \quad \mathbf{e}_2' = [0.999 \quad -0.040]$$

Eigenvalue - eigenvector pairs for $\boldsymbol{\rho}$:

$$\lambda_1 = 1 + \rho = 1.4, \quad \mathbf{e}_1' = [0.707 \quad 0.707]$$

$$\lambda_2 = 1 - \rho = 0.6, \quad \mathbf{e}_2' = [0.707 \quad -0.707]$$

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Example 8.2

Principal components for $\boldsymbol{\Sigma}$: $\frac{\lambda_1}{\lambda_1 + \lambda_2} = 0.992$

$$Y_1 = 0.040X_1 + 0.999X_2$$

$$Y_2 = 0.999X_1 - 0.040X_2$$

Principal components for $\boldsymbol{\rho}$: $\frac{\lambda_1}{p} = 0.7$

$$Y_1 = 0.707Z_1 + 0.707Z_2 = 0.707(X_1 - \mu_1) + 0.707(X_2 - \mu_2)$$

$$Y_2 = 0.707Z_1 - 0.707Z_2 = 0.707(X_1 - \mu_1) - 0.707(X_2 - \mu_2)$$

$$\rho_{Y_1, Z_1} = e_{11} \sqrt{\lambda_1} = 0.837, \quad \rho_{Y_1, Z_2} = e_{12} \sqrt{\lambda_1} = 0.837$$

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Principal Components for Diagonal Covariance Matrix

$$\boldsymbol{\Sigma} = \begin{bmatrix} \sigma_{11} & 0 & \dots & 0 \\ 0 & \sigma_{22} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \sigma_{pp} \end{bmatrix}, \quad \mathbf{e}_i' = [0 \quad \dots \quad 0 \quad 1 \quad 0 \quad \dots \quad 0]$$

$$\boldsymbol{\Sigma} \mathbf{e}_i = \sigma_{ii} \mathbf{e}_i, \quad Y_i = \mathbf{e}_i' \mathbf{X} = X_i$$

$$\boldsymbol{\rho} = \mathbf{I}, \quad \boldsymbol{\rho} \mathbf{e}_i = 1 \mathbf{e}_i, \quad Y_i = \mathbf{e}_i' \mathbf{Z} = Z_i$$

\mathbf{X} : $N_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, constant density ellipsoid is
a right ellipsoid for \mathbf{X}
and a sphere for \mathbf{Z}

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Principal Components for a Special Covariance Matrix

$$\boldsymbol{\Sigma} = \begin{bmatrix} \sigma^2 & \rho\sigma^2 & \dots & \rho\sigma^2 \\ \rho\sigma^2 & \sigma^2 & \dots & \rho\sigma^2 \\ \vdots & \vdots & \ddots & \vdots \\ \rho\sigma^2 & \rho\sigma^2 & \dots & \sigma^2 \end{bmatrix}, \quad \boldsymbol{\rho} = \begin{bmatrix} 1 & \rho & \dots & \rho \\ \rho & 1 & \dots & \rho \\ \vdots & \vdots & \ddots & \vdots \\ \rho & \rho & \dots & 1 \end{bmatrix}$$

$$\lambda_1 = 1 + (p-1)\rho, \quad \mathbf{e}_1' = \left[\frac{1}{\sqrt{p}} \quad \frac{1}{\sqrt{p}} \quad \dots \quad \frac{1}{\sqrt{p}} \right]$$

$$\lambda_2 = \lambda_3 = \dots = \lambda_p = 1 - \rho$$

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Principal Components for a Special Covariance Matrix

$$\mathbf{e}_i' = \begin{bmatrix} \frac{1}{\sqrt{(i-1)i}} & \cdots & \frac{1}{\sqrt{(i-1)i}} & \frac{-(i-1)}{\sqrt{(i-1)i}} & 0 & \cdots & 0 \end{bmatrix}$$

$$i = 2, \dots, p$$

$$Y_1 = \mathbf{e}_1' \mathbf{Z} = \frac{1}{\sqrt{p}} \sum_{i=1}^p Z_i, \quad \frac{\lambda_1}{p} = \rho + \frac{1-\rho}{p}$$

the last $p-1$ components collectively contribute very little to the total variance and can be neglected when ρ is near 1

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Questions

- What are the sample principal components?
- How to compute the sample principal components?
- How to decide the number of principal components required?
- What is the geometric interpretation of the sample principal components?

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Questions

- How to compute the sample principal components for standardized random vectors?
- What does it mean for an unusually small value for the last eigenvalue from either the sample covariance or correlation matrix?

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Sample Principal Components

$\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n$: n independent drawings from some p -dimensional population with mean $\boldsymbol{\mu}$ and covariance matrix $\boldsymbol{\Sigma}$

sample mean $\bar{\mathbf{x}}$, sample covariance matrix \mathbf{S}

first sample principal component $\mathbf{a}_1' \mathbf{x}_j$:

$$\max_{\mathbf{a}_1} \mathbf{a}_1' \mathbf{S} \mathbf{a}_1 \quad \text{subject to } \mathbf{a}_1' \mathbf{a}_1 = 1$$

i th sample principal component $\mathbf{a}_i' \mathbf{x}_j$:

$$\max_{\mathbf{a}_i} \mathbf{a}_i' \mathbf{S} \mathbf{a}_i \quad \text{subject to } \mathbf{a}_i' \mathbf{a}_i = 1 \quad \text{and} \quad \mathbf{a}_i' \mathbf{S} \mathbf{a}_k = 0$$

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Sample Principal Components

$\mathbf{S} = \{s_{ik}\}$ is with eigenvalue - eigenvector pairs

$$(\hat{\lambda}_i, \hat{\mathbf{e}}_i), \quad i, k = 1, 2, \dots, p$$

i th sample principal component of observation \mathbf{x} :

$$\hat{y}_i = \hat{\mathbf{e}}_i' \mathbf{x} = \hat{e}_{i1}x_1 + \hat{e}_{i2}x_2 + \cdots + \hat{e}_{ip}x_p$$

$$\hat{\lambda}_1 \geq \hat{\lambda}_2 \geq \cdots \geq \hat{\lambda}_p \geq 0$$

$$\text{sample variance}(\hat{y}_k) = \hat{\lambda}_k$$

$$\text{sample covariance}(\hat{y}_i, \hat{y}_k) = 0, \quad i \neq k$$

$$\text{Total sample variance} = \sum_{i=1}^p s_{ii} = \sum_{i=1}^p \hat{\lambda}_i, \quad \hat{r}_{\hat{y}_i, \hat{y}_k} = \frac{\hat{e}_{ik} \sqrt{\hat{\lambda}_i}}{\sqrt{s_{kk}}} \quad 36$$

Example 8.3

Socioeconomic variables for 61 tracts in Madison, Wisconsin.

X_1 : total population (thousands)

X_2 : professional degree (percent)

X_3 : employed age over 16 (percent)

X_4 : government employment (percent)

X_5 : median home value (\$10,000s)

$$\bar{\mathbf{x}}' = [4.47 \quad 3.96 \quad 71.42 \quad 26.91 \quad 1.64]$$

$$\mathbf{S} = \begin{bmatrix} 3.397 & & & & \\ -1.102 & 9.673 & & & \\ 4.306 & -1.513 & 55.626 & & \\ -2.078 & 10.953 & -28.937 & 89.067 & \\ 0.027 & 1.203 & -0.044 & 0.957 & 0.319 \end{bmatrix}$$

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Example 8.3

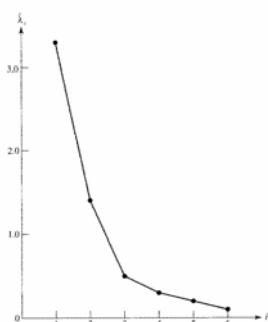
Coefficients for the Principal Components

(Correlation Coefficients in Parentheses)

Variable	$\hat{\mathbf{e}}_1 (r_{1, \hat{\mathbf{e}}_1})$	$\hat{\mathbf{e}}_2 (r_{2, \hat{\mathbf{e}}_2})$	$\hat{\mathbf{e}}_3$	$\hat{\mathbf{e}}_4$	$\hat{\mathbf{e}}_5$
Total population	-0.039(-.22)	0.071(.24)	0.188	0.977	-0.058
Profession	0.105(.35)	0.130(.26)	-0.961	0.171	-0.199
Employment (%)	-0.492(-.68)	0.864(.73)	0.046	-0.091	0.005
Government employment (%)	0.863(.95)	0.480(.32)	0.153	-0.030	0.007
Median home value	0.009(.16)	0.015(.17)	-0.125	0.082	0.989
Variance ($\hat{\lambda}_i$):	107.02	39.67	8.37	2.87	0.15
Cumulative percentage of total variance	67.7	92.8	98.1	99.9	1.000

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Scree Plot to Determine Number of Principal Components



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Example 8.4: Pained Turtles

natural logarithms of the measured carapace

length, width, and weight of 24 male pained turtles

sample mean vector :

$$\bar{\mathbf{x}} = [4.725 \quad 4.478 \quad 3.703]$$

sample covariance matrix

$$\mathbf{S} = 10^{-3} \begin{bmatrix} 11.072 & 8.019 & 8.160 \\ 8.019 & 6.417 & 6.005 \\ 8.160 & 6.005 & 6.773 \end{bmatrix}$$

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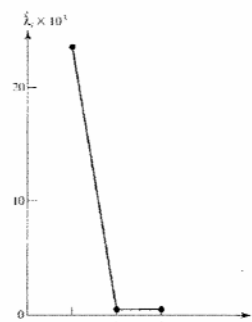
Example 8.4

COEFFICIENTS FOR PRINCIPAL COMPONENTS
(Correlation Coefficients in Parentheses)

Variable	$\hat{\mathbf{e}}_1 (r_{1, \hat{\mathbf{e}}_1})$	$\hat{\mathbf{e}}_2$	$\hat{\mathbf{e}}_3$
ln (length)	.683 (.99)	-.159	-.713
ln (width)	.510 (.97)	-.594	.622
ln (height)	.523 (.97)	.788	.324
Variance ($\hat{\lambda}_i$):	23.30×10^{-3}	$.60 \times 10^{-3}$	$.36 \times 10^{-3}$
Cumulative percentage of total variance	96.1	98.5	100

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Example 8.4: Scree Plot



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Example 8.4: Principal Component

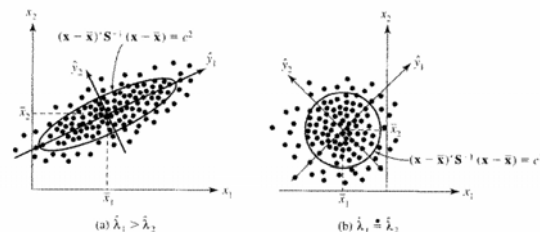
- One dominant principal component
 - Explains 96% of the total variance

Interpretation

$$\begin{aligned}\hat{y}_1 &= 0.683 \ln(\text{length}) + 0.510 \ln(\text{width}) + 0.523 \ln(\text{height}) \\ &= \ln[(\text{length})^{0.683} (\text{width})^{0.510} (\text{height})^{0.523}] \\ &= \ln(\text{volume of a box with adjusted dimension})\end{aligned}$$

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Geometric Interpretation



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Standardized Variables

$$z_{ji} = \frac{x_{ji} - \bar{x}_i}{\sqrt{s_{ii}}}, \quad i = 1, 2, \dots, p, \quad \mathbf{Z} = \{z_{ji}\}$$

$$\mathbf{z}_j = \mathbf{D}^{-1/2}(\mathbf{x}_j - \bar{\mathbf{x}}), \quad \bar{\mathbf{z}} = \frac{1}{n} \mathbf{Z}' \mathbf{1} = \mathbf{0}$$

$$\mathbf{S}_z = \frac{1}{n-1} \mathbf{Z}' \mathbf{Z} = \begin{bmatrix} 1 & \frac{s_{12}}{\sqrt{s_{11}}\sqrt{s_{22}}} & \dots & \frac{s_{1p}}{\sqrt{s_{11}}\sqrt{s_{pp}}} \\ \frac{s_{21}}{\sqrt{s_{11}}\sqrt{s_{22}}} & 1 & \dots & \frac{s_{2p}}{\sqrt{s_{22}}\sqrt{s_{pp}}} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{s_{p1}}{\sqrt{s_{11}}\sqrt{s_{pp}}} & \frac{s_{p2}}{\sqrt{s_{22}}\sqrt{s_{pp}}} & \dots & 1 \end{bmatrix} = \mathbf{R}$$

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Principal Components

$\mathbf{z}_1, \mathbf{z}_2, \dots, \mathbf{z}_n$ are standardized observations
with sample covariance matrix \mathbf{R}

$(\hat{\lambda}_i, \hat{\mathbf{e}}_i)$: eigenvalue-eigenvector pairs of \mathbf{R}

$$\hat{\lambda}_1 \geq \hat{\lambda}_2 \geq \dots \geq \hat{\lambda}_p \geq 0$$

The i th principal component of \mathbf{z} :

$$\hat{y}_i = \hat{\mathbf{e}}_i' \mathbf{z}, \quad i = 1, 2, \dots, p$$

sample variance $(\hat{y}_i) = \hat{\lambda}_i$, sample covariance $(\hat{y}_i, \hat{y}_k) = 0, i \neq k$

total sample variance = $\text{tr}(\mathbf{R}) = p$

$$r_{y_i, z_k} = \hat{e}_{ik} \sqrt{\hat{\lambda}_i}, \quad i, k = 1, 2, \dots, p$$

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Proportion of Total Variance due to the k th Principal Component

$$\left(\begin{array}{l} \text{Proportion of (standardized)} \\ \text{sample variance} \\ \text{due to the } k\text{th sample principal} \\ \text{component} \end{array} \right) = \frac{\hat{\lambda}_k}{p}, \quad k = 1, 2, \dots, p$$

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Example 8.5: Stocks Data

- Weekly rates of return for five stocks
 - X_1 : JP Morgan
 - X_2 : Citibank
 - X_3 : Wells Fargo
 - X_4 : Royal Dutch Shell
 - X_5 : ExxonMobil

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Example 8.5

$$\bar{\mathbf{x}}' = [0.0011 \quad 0.0007 \quad 0.0016 \quad 0.0040 \quad 0.0040]$$

$$\mathbf{R} = \begin{bmatrix} 1 & & & & \\ 0.632 & 1 & & & \\ 0.511 & 0.574 & 1 & & \\ 0.115 & 0.322 & 0.183 & 1 & \\ 0.155 & 0.213 & 0.146 & 0.683 & 1 \end{bmatrix}$$

$$\hat{\lambda}_1 = 2.437, \quad \hat{\mathbf{e}}_1' = [0.469 \quad 0.532 \quad 0.465 \quad 0.387 \quad 0.361]$$

$$\hat{\lambda}_2 = 1.407, \quad \hat{\mathbf{e}}_2' = [-0.368 \quad -0.236 \quad -0.315 \quad 0.585 \quad 0.606]$$

$$\hat{\lambda}_3 = 0.501, \quad \hat{\mathbf{e}}_3' = [-0.604 \quad -0.136 \quad 0.772 \quad 0.093 \quad -0.109]$$

$$\hat{\lambda}_4 = 0.400, \quad \hat{\mathbf{e}}_4' = [0.363 \quad -0.629 \quad 0.289 \quad -0.381 \quad 0.493]$$

$$\hat{\lambda}_5 = 0.255, \quad \hat{\mathbf{e}}_5' = [0.384 \quad -0.496 \quad 0.071 \quad 0.595 \quad -0.498]$$

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Example 8.5

First two principal components :

$$\hat{y}_1 = \hat{\mathbf{e}}_1' \mathbf{z} = 0.469z_1 + 0.532z_2 + 0.465z_3 + 0.387z_4 + 0.361z_5$$

$$\hat{y}_2 = \hat{\mathbf{e}}_2' \mathbf{z} = -0.368z_1 - 0.236z_2 - 0.315z_3 + 0.585z_4 + 0.606z_5$$

$$\frac{\hat{\lambda}_1 + \hat{\lambda}_2}{p} = 77\%$$

\hat{y}_1 : roughly equally weighted sum (index) of the five stocks
(general stock - market component, or, market component)

\hat{y}_2 : contrast banking stocks and the oil stocks
(industry component)

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Example 8.6

- Body weight (in grams) for $n=150$ female mice were obtained after the birth of their first 4 litters

$$\bar{\mathbf{x}}' = [39.88 \quad 45.08 \quad 48.11 \quad 49.95]$$

$$\mathbf{R} = \begin{bmatrix} 1 & & & \\ 0.7501 & 1 & & \\ 0.6329 & 0.6925 & 1 & \\ 0.6363 & 0.7386 & 0.6625 & 1 \end{bmatrix}$$

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Example 8.6

$$\hat{\lambda}_1 = 3.085, \quad \hat{\lambda}_2 = 0.382, \quad \hat{\lambda}_3 = 0.342, \quad \hat{\lambda}_4 = 0.217$$

$$\hat{\lambda}_1 \approx 1 + (p-1)\bar{r} = 1 + (4-1) \times 0.6854 = 3.056$$

$$\hat{\lambda}_2 \approx \hat{\lambda}_3 \approx \hat{\lambda}_4 \ll \hat{\lambda}_1$$

$$\hat{y}_1 = \hat{\mathbf{e}}_1' \mathbf{z} = 0.49z_1 + 0.52z_2 + 0.49z_3 + 0.50z_4$$

$$\frac{\hat{\lambda}_1}{p} = 0.76$$

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Comment

- An unusually small value for the last eigenvalue from either the sample covariance or correlation matrix can indicate an unnoticed linear dependency of the data set
- One or more of the variables is redundant and should be deleted
- Example: $x_4 = x_1 + x_2 + x_3$

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Outline

- Introduction
- Popular Principal Components
- Summarizing Sample Variation by Principal Components
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- Large Sample Inferences
- Monitoring Quality with Principal Components

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Questions

- Why to check the normality of the first few principal components?
- How to pinpoint suspect observation?

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Check Normality and Suspect Observations

- Construct scatter diagram for pairs of the first few principal components
- Make $Q-Q$ plots from the sample values generated by each principal component
- Construct scatter diagram and $Q-Q$ plots for the last few principal components

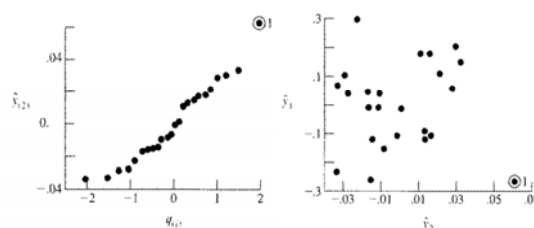
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Example 8.7: Turtle Data

$$\begin{aligned}\hat{y}_1 &= 0.683(x_1 - 4.725) + 0.510(x_2 - 4.478) \\ &\quad + 0.523(x_3 - 3.703) \\ \hat{y}_2 &= -0.159(x_1 - 4.725) - 0.594(x_2 - 4.478) \\ &\quad + 0.788(x_3 - 3.703) \\ \hat{y}_3 &= -0.713(x_1 - 4.725) + 0.622(x_2 - 4.478) \\ &\quad + 0.324(x_3 - 3.703)\end{aligned}$$

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Example 8.7



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Questions

- What are the large sample distribution for eigenvalues and eigenvectors?
- How to determine the confidence interval for an eigenvalue?
- What is the approximate distribution for estimated eigenvectors?
- How to test for equal correlation structure?

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Large Sample Distribution for Eigenvalues and Eigenvectors

\mathbf{S} is with eigen values $\hat{\lambda}' = [\hat{\lambda}_1 \ \cdots \ \hat{\lambda}_p]$ and

eigenvectors $\hat{\mathbf{e}}_1, \hat{\mathbf{e}}_2, \dots, \hat{\mathbf{e}}_p$

Let $\mathbf{\Lambda} = \text{diag}\{\lambda_1, \dots, \lambda_p\}$, λ_i 's are eigenvalues of $\mathbf{\Sigma}$

$\Rightarrow \sqrt{n}(\hat{\lambda} - \lambda)$: approximately $N_p(\mathbf{0}, 2\mathbf{\Lambda}^2)$

Let $\mathbf{E}_i = \lambda_i \sum_{\substack{k=1 \\ k \neq i}}^p \frac{\lambda_k}{(\lambda_k - \lambda_i)^2} \mathbf{e}_k \mathbf{e}_k'$

$\Rightarrow \sqrt{n}(\hat{\mathbf{e}}_i - \mathbf{e}_i)$: approximately $N_p(\mathbf{0}, \mathbf{E}_i)$

$\hat{\lambda}_i$ is independent of the elements of associated $\hat{\mathbf{e}}_i$ 61

Confidence Interval for λ_i

$\hat{\lambda}_i : N(\lambda_i, 2\lambda_i^2/n)$ for n large

$$P\left[\frac{|\hat{\lambda}_i - \lambda_i|}{\lambda_i \sqrt{\frac{2}{n}}} \leq z\left(\frac{\alpha}{2}\right)\right] = 1 - \alpha$$

100(1 - α)% confidence interval for λ_i :

$$\frac{\hat{\lambda}_i}{1 + z(\alpha/2)\sqrt{2/n}} \leq \lambda_i \leq \frac{\hat{\lambda}_i}{1 - z(\alpha/2)\sqrt{2/n}}$$

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Approximate Distribution of Estimated Eigenvectors

$\sqrt{n}(\hat{\mathbf{e}}_i - \mathbf{e}_i)$: approximate $N_p(\mathbf{0}, \mathbf{E}_i)$

\mathbf{E}_i can be approximated by

$$\hat{\mathbf{E}}_i = \hat{\lambda}_i \sum_{\substack{k=1 \\ k \neq i}}^p \frac{\hat{\lambda}_k}{(\hat{\lambda}_k - \hat{\lambda}_i)^2} \hat{\mathbf{e}}_k \hat{\mathbf{e}}_k'$$

$$\hat{e}_{ik} : N(e_{ik}, \hat{E}_{i,kk}/n)$$

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Example 8.8

Stock price data : $N_5(\boldsymbol{\mu}, \mathbf{\Sigma})$

$\mathbf{\Sigma}$ has distinct eigenvalues $\lambda_1 > \lambda_2 > \dots > \lambda_5 > 0$

$n = 103$ large

$\hat{\lambda}_1 = 0.0014$, $z(0.025) = 1.96$

95% confidence interval

$$\frac{0.0014}{1 + 1.96\sqrt{2/103}} \leq \lambda_1 \leq \frac{0.0014}{1 - 1.96\sqrt{2/103}}, \text{ or } 0.0011 \leq \lambda_1 \leq 0.0019$$

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Testing for Equal Correlation

$$H_0 : \boldsymbol{\rho} = \boldsymbol{\rho}_0 = \begin{bmatrix} 1 & \rho & \cdots & \rho \\ \rho & 1 & \cdots & \rho \\ \vdots & \vdots & \ddots & \vdots \\ \rho & \rho & \cdots & 1 \end{bmatrix}, \quad H_1 : \boldsymbol{\rho} \neq \boldsymbol{\rho}_0$$

$$\bar{r}_k = \frac{1}{p-1} \sum_{i=1}^p r_{ik}, \quad \bar{r} = \frac{2}{p(p-2)} \sum_k \sum_{i < k} r_{ik}, \quad \hat{\gamma} = \frac{(p-1)^2 [1 - (1 - \bar{r})^2]}{p - (p-2)(1 - \bar{r})^2}$$

Reject H_0 in favor of H_1 if

$$T = \frac{(n-1)}{(1 - \bar{r})^2} \left[\sum_k \sum_{i < k} (r_{ik} - \bar{r})^2 - \hat{\gamma} \sum_{k=1}^p (\bar{r}_k - \bar{r})^2 \right] > \chi_{(p+1)(p-2)/2}^2(\alpha)$$

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Example 8.9

Example 8.6, female mice data $\mathbf{R} = \begin{bmatrix} 1 & & & & \\ 0.7501 & 1 & & & \\ 0.6329 & 0.6925 & 1 & & \\ 0.6363 & 0.7386 & 0.6625 & 1 & \\ & & & & \end{bmatrix}$

$$\bar{r}_1 = 0.6731, \bar{r}_2 = 0.7271, \bar{r}_3 = 0.6626, \bar{r}_4 = 0.6791, \bar{r} = 0.6855$$

$$\sum_k \sum_{i < k} (r_{ik} - \bar{r})^2 = 0.01277, \sum_{k=1}^4 (\bar{r}_k - \bar{r})^2 = 0.00245, \hat{\gamma} = 2.1329$$

$$T = \frac{(150-1)}{(1 - 0.6855)^2} [0.01277 - (2.1329)(0.00245)] = 11.4$$

$$> \chi_{(4+1)(4-2)/2}^2(0.05) = 11.07$$

The evidence against H_0 is strong, but not overwhelming

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Questions

- How to monitor a stable process using the first two principal components?
- How to monitor a stable process using the T^2 chart from the principal components?
- How to control future values by principal components?

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Questions

- Why avoiding Computation with Small Eigenvalues?

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Monitoring Stable Process: Part 1

The values of the first two principal components should be stable for a process stable over time

Construct the quality ellipse for the first two principal components when n large :

$$\frac{\hat{y}_1^2}{\hat{\lambda}_1} + \frac{\hat{y}_2^2}{\hat{\lambda}_2} \leq \chi^2_2(\alpha)$$

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Example 8.10 Police Department Data

Variable	\hat{e}_1	\hat{e}_2	\hat{e}_3	\hat{e}_4	\hat{e}_5
Appearances overtime (x_1)	.046	-.048	.629	-.643	.432
Extraordinary event (x_2)	.039	.985	-.077	-.151	-.007
Holdover hours (x_3)	-.658	.107	.582	.250	-.392
COA hours (x_4)	.734	.069	.503	.397	-.213
Meeting hours (x_5)	-.155	.107	.081	.586	.784
$\hat{\lambda}_i$	2,770,226	1,429,206	628,129	221,138	99,824

*First two sample components explain 82% of the total variance

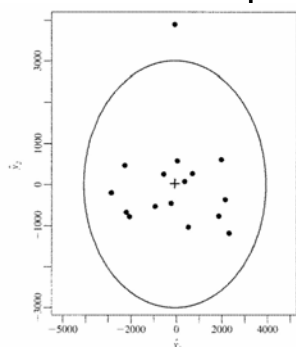
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Example 8.10: Principal Components

Period	\hat{y}_{j1}	\hat{y}_{j2}	\hat{y}_{j3}	\hat{y}_{j4}	\hat{y}_{j5}
1	2044.9	588.2	425.8	-189.1	-209.8
2	-2143.7	-686.2	883.6	-565.9	-441.5
3	-177.8	-464.6	707.5	736.3	38.2
4	-2186.2	450.5	-184.0	443.7	-325.3
5	-878.6	-545.7	115.7	296.4	437.5
6	563.2	-1045.4	281.2	620.5	142.7
7	403.1	66.8	340.6	-135.5	521.2
8	-1988.9	-801.8	-1437.3	-148.8	61.6
9	132.8	563.7	125.3	68.2	611.5
10	-2787.3	-213.4	7.8	169.4	-202.3
11	283.4	3936.9	-0.9	276.2	-159.6
12	761.6	256.0	-2153.6	-418.8	28.2
13	-498.3	244.7	966.5	-1142.3	182.6
14	2366.2	-1193.7	-165.5	270.6	-344.9
15	1917.8	-782.0	-82.9	-196.8	-89.9
16	2187.7	-373.8	170.1	-84.1	-250.2

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Example 8.10:
95% Control Ellipse



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Monitoring Stable Process: Part 2

$$\mathbf{X} : N_p(\boldsymbol{\mu}, \boldsymbol{\Sigma}), \quad \mathbf{E} = [\mathbf{e}_1 \quad \mathbf{e}_2 \quad \dots \quad \mathbf{e}_p]$$

$$\mathbf{X} - \boldsymbol{\mu} = \sum_{i=1}^p (\mathbf{X} - \boldsymbol{\mu})' \mathbf{e}_i \mathbf{e}_i = \sum_{i=1}^p Y_i \mathbf{e}_i$$

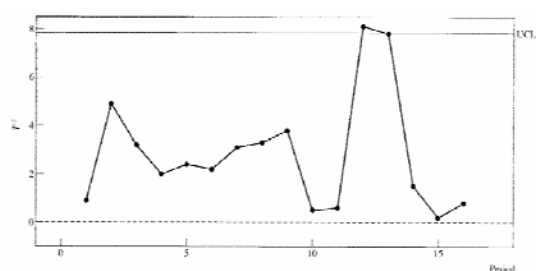
$$\mathbf{E}'(\mathbf{X} - \boldsymbol{\mu} - Y_1 \mathbf{e}_1 - Y_2 \mathbf{e}_2) = \begin{bmatrix} 0 & 0 & Y_3 & \dots & Y_p \end{bmatrix} = \begin{bmatrix} 0 & 0 & \mathbf{Y}_{(2)} \end{bmatrix}$$

$$\mathbf{Y}_{(2)}' \boldsymbol{\Sigma}_{\mathbf{Y}_{(2)}, \mathbf{Y}_{(2)}}^{-1} \mathbf{Y}_{(2)} = \frac{Y_3^2}{\lambda_3} + \frac{Y_4^2}{\lambda_4} + \dots + \frac{Y_p^2}{\lambda_p} : \chi_{p-2}^2$$

$$T_j^2 = \frac{\hat{y}_{j3}^2}{\hat{\lambda}_3} + \frac{\hat{y}_{j4}^2}{\hat{\lambda}_4} + \dots + \frac{\hat{y}_{jp}^2}{\hat{\lambda}_p}, \quad \text{UCL} = \chi_{p-2}^2(\alpha)$$

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Example 8.11
 T^2 Chart for Unexplained Data



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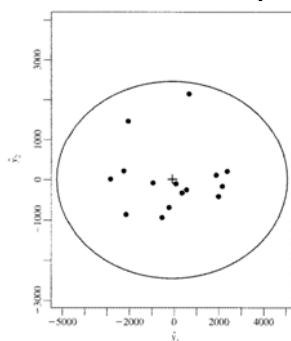
Example 8.12
Control Ellipse for Future Values

	$\hat{\mathbf{e}}_1$	$\hat{\mathbf{e}}_2$	$\hat{\mathbf{e}}_3$	$\hat{\mathbf{e}}_4$	$\hat{\mathbf{e}}_5$
Appearances overtime (x_1)	.049	.629	.304	.479	.530
Extraordinary event (x_2)	.007	-.078	.939	-.260	-.212
Holdover hours (x_3)	-.662	.582	-.089	-.158	-.437
COA hours (x_4)	.731	.503	-.123	-.336	-.291
Meeting hours (x_5)	-.159	.081	-.058	-.752	.632
$\hat{\lambda}_i$	2,964,749.9	672,995.1	396,596.5	194,401.0	92,760.3

*Example 8.10 data after dropping out-of-control case

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Example 8.12
99% Prediction Ellipse



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Avoiding Computation with Small Eigenvalues

$$\begin{aligned} d_{Uj}^2 &= (\bar{\mathbf{x}}_j - \bar{\mathbf{x}} - \hat{y}_{j1} \hat{\mathbf{e}}_1 - \hat{y}_{j2} \hat{\mathbf{e}}_2) (\bar{\mathbf{x}}_j - \bar{\mathbf{x}} - \hat{y}_{j1} \hat{\mathbf{e}}_1 - \hat{y}_{j2} \hat{\mathbf{e}}_2)' \\ &= (\bar{\mathbf{x}}_j - \bar{\mathbf{x}} - \hat{y}_{j1} \hat{\mathbf{e}}_1 - \hat{y}_{j2} \hat{\mathbf{e}}_2)' \hat{\mathbf{E}} \hat{\mathbf{E}}' (\bar{\mathbf{x}}_j - \bar{\mathbf{x}} - \hat{y}_{j1} \hat{\mathbf{e}}_1 - \hat{y}_{j2} \hat{\mathbf{e}}_2) \\ &= \sum_{k=3}^p \hat{y}_{jk}^2 : \text{approximate } c\chi_v^2 \end{aligned}$$

$$\bar{d}_U^2 = \frac{1}{n} \sum_{j=1}^n d_{Uj}^2 = c\nu, \quad s_{d^2}^2 = \frac{1}{n-1} \sum_{j=1}^n (d_{Uj}^2 - \bar{d}_U^2)^2 = 2c^2\nu$$

$$c = \frac{s_{d^2}^2}{2\bar{d}_U^2}, \quad \nu = 2 \frac{(\bar{d}_U^2)^2}{s_{d^2}^2}$$

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